

Do 8 of the following 10 problems.

1. Let  $X$  be a discrete random variable that can only assume the values  $0, 1, 2, \dots$ . Prove that

$$E(X) = \sum_{k=0}^{\infty} P\{X > k\}$$

2. Let  $X_1, \dots, X_n$  be a random sample from the density

$$f(x; \theta) = \begin{cases} \frac{\log \theta}{\theta - 1} \theta^x & , 0 < x < 1, \theta > 1 \\ 0 & , \text{elsewhere} \end{cases}$$

(You may find it helpful to realize that  $\theta^x = e^{x \log \theta}$ .)

- (a) Compute  $E(X_i)$ ,  $i = 1, \dots, n$ .

- (b) Find a sufficient statistic for  $\theta$ .

- (c) Compute the Fisher information in the random sample.

3. Let  $X$  and  $Y$  be independent random variables with p.d.f.'s  $f_x(x) = \begin{cases} \frac{1}{\theta} e^{-x/\theta} & , x > 0, \theta > 0 \\ 0 & , \text{elsewhere} \end{cases}$

$$\text{and } f_y(y) = \begin{cases} \frac{1}{\alpha} e^{-y/\alpha} & , y > 0, \alpha > 0 \\ 0 & , \text{elsewhere} \end{cases} \text{ respectively.}$$

- (a) Find  $P(X \leq Y)$

- (b) Find  $P(X \leq t, X \leq Y)$

4. Let  $X_1, \dots, X_n$  be i.i.d. random variables with c.d.f.  $F(x; \theta)$  determined by the odds ratio

$$\frac{F(x; \theta)}{1 - F(x; \theta)} = R(x; \theta).$$

Find the method of moments estimator of  $\theta$  if

- (a)  $R(x; \theta) = e^{(x-\theta)}$ ,  $x > \theta$ ,  $-\infty < \theta < \infty$

- (b)  $R(x; \theta) = (x - \theta)^2$ ,  $x > \theta$ ,  $-\infty < \theta < \infty$

$$\left( \text{Hint: } \int_0^{\infty} \frac{dt}{(1+t^2)^2} = \frac{\pi}{4} \right)$$

5. Let  $X_1, \dots, X_n$  denote a random sample from the uniform distribution on the interval  $(0, \theta)$ .

- (a) Find an unbiased estimator for  $\theta$ , based on the sample mean  $\bar{x}$ . Denote it by  $\hat{\theta}_1$ .

- (b) Find an unbiased estimator for  $\theta$ , based on the  $n^{\text{th}}$  order statistic  $X_{(n)} = \max(X_1, \dots, X_n)$ . Denote it by  $\hat{\theta}_2$ .

- (c) Define the efficiency of  $\hat{\theta}_1$  relative to  $\hat{\theta}_2$  by the ratio  $\text{Var}(\hat{\theta}_2) / \text{Var}(\hat{\theta}_1) = \text{Eff}(\hat{\theta}_1, \hat{\theta}_2)$ . Find  $\text{Eff}(\hat{\theta}_1, \hat{\theta}_2)$ .

6. Suppose  $n$  integers are drawn from  $\{1, 2, \dots, N\}$  at random and with replacement.

(a) Find the method of moments estimator,  $\hat{N}_1$ , of  $N$ .

(b) Find  $E(\hat{N}_1)$ .

(c) Find  $\text{Var}(\hat{N}_1)$ .

(d) Find the maximum-likelihood estimator,  $\hat{N}_2$  of  $N$ .

7. Suppose that eight observations  $X_1, \dots, X_8$  are drawn at random from a distribution with the following p.d.f.

$$f(x; \theta) = \begin{cases} \theta x^{\theta-1} & , 0 < x < 1, \theta > 0 \\ 0 & , \text{elsewhere} \end{cases}$$

Show that the UMP test of

$$H_0: \theta = 1 \text{ against } H_1: \theta > 1$$

at the  $\alpha$  level of significance is given by the rejection region  $C = \{(X_1, \dots, X_8) : \sum_{i=1}^8 \ln X_i \leq c\}$ . Explain how the constant  $c$  is found.

8. Let  $X_1, \dots, X_n$  be a random sample from the distribution that has the p.d.f.

$$f(x; \theta) = \begin{cases} \theta^x e^{-\theta} / x! & , x = 0, 1, 2, \dots; \theta > 0 \\ 0 & , \text{elsewhere} \end{cases}$$

(a) Show that  $T = \sum_{i=1}^n X_i$  is a complete sufficient statistic for  $\theta$ .

(b) Construct the best unbiased estimator for  $\theta$ . Does the variance of this estimator attain the Rao-Cramér bound?

(c) Construct the best unbiased estimator for the parameter  $\theta^2$ .

9. Let  $X_1, \dots, X_n$  be a random sample of size  $n$  from a gamma distribution with p.d.f.

$$f(x; \alpha, \beta) = \begin{cases} \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta} & , x > 0, \alpha > 0, \beta > 0 \\ 0 & , \text{elsewhere} \end{cases}$$

(a) Show that  $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$  and  $\tilde{X} = (\prod_{i=1}^n X_i)^{1/n}$  are jointly complete and sufficient for  $\alpha$  and  $\beta$ .

(b) For  $\alpha = 2$ , find the UMVUE of  $\beta$ . Does the variance of this estimator attain the Rao-Cramér bound?

(c) Show that the distribution of  $T = \frac{\bar{X}}{\tilde{X}}$  does not depend on  $\beta$ . Is  $T$  independent of  $\bar{X}$ ? Explain.

10. Let  $X_1, \dots, X_n$  be a random sample from a binomial distribution with p.d.f.

$$f(x; p) = \begin{cases} \binom{m}{x} p^x (1-p)^{m-x} & , x = 0, 1, \dots, m, 0 \leq p \leq 1 \\ 0 & , \text{elsewhere} \end{cases}$$

(a) Find the UMVUE of  $p$ .

(b) Show that the estimator in part (a) converges in probability to  $p$  as  $n \rightarrow \infty$ .