

Preliminary Exam: Probability and Statistics
2004

Work all 9 problems. Begin each problem on a new page, using one side of the sheet. Tables of the normal and chi-square distributions are attached. Throughout “p.d.f.” means “probability density function,” \mathbb{R} is the real line, and $\mathbf{1}_A$ the indicator function of the set A (i.e. $\mathbf{1}_A(x) = 1$ if $x \in A, 0$ otherwise).

1. Let X and Y be nonnegative random variables of continuous type with p.d.f.’s f and g respectively. These densities are supposed to be continuous on \mathbb{R} with respective cumulative distribution functions F and G satisfying $\lim_{x \rightarrow \infty} x\{1 - F(x)\} = 0$ and $\lim_{x \rightarrow \infty} x\{1 - G(x)\} = 0$.

a. Prove that $E(X) = \int_0^\infty \{1 - F(x)\}dx$.

b. If $P\{X > x\} \geq P\{Y > x\}$ for all $x \in \mathbb{R}$, prove that $E(X) \geq E(Y)$.

2. Let X_1, \dots, X_n be independent and identically distributed random variables with common density

$$f_\theta(x) = \theta x^{-\theta-1} \mathbf{1}_{(1, \infty)}(x), x \in \mathbb{R},$$

$\theta > 0$. Define $S_n = (\prod_{i=1}^n X_i)^{1/n}$.

a. Compute $E_\theta(\log S_n)$ and $\text{Var}_\theta(\log S_n)$.

b. Find the limiting distribution of $\log S_n$.

c. Choose $\theta = 10, n = 100$, and approximate $P\{S_n > e^{0.11}\}$.

3. Let the random variables X and Y have the joint p.d.f.

$$f(x, y) = \begin{cases} 8xy, & \text{for } 0 < x < y < 1, \\ 0, & \text{elsewhere.} \end{cases}$$

a. Determine the marginal p.d.f.’s.

b. Compute $E(Y), \text{Var}(Y)$.

c. Find the conditional p.d.f. of Y given $X = x$.

d. Find the conditional expectation of Y given $X = x$.

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4. Let X be a random sample of size 1 from the p.d.f.

$$f_{\theta}(x) = \frac{1}{2\theta} \mathbf{1}_{(-\theta, \theta)}(x), x \in \mathbb{R},$$

$\theta > 0$.

- Compute $E_{\theta}(\sin X)$.
- Is X complete for $\theta > 0$? (Explain your answer.)
- Let X_1, \dots, X_n be a random sample of size n from the above p.d.f. Find a real valued random variable that is a sufficient statistic for $\theta > 0$.

5. Let X_1, \dots, X_n be a random sample of size n from the Poisson (θ) - distribution.

- Explain why $T = \sum_{i=1}^n X_i$ is a sufficient and complete statistic for $\theta > 0$.
- Find the exact probability distribution of T using the moment generating function technique.
- Find the uniform minimum variance unbiased estimator of $h(\theta) = \theta^2, \theta > 0$.
(Hint: you may use that $\sum_{k=2}^{\infty} k(k-1)e^{-n\theta} \frac{(n\theta)^k}{k!} = (n\theta)^2$.)

6. Let X_1, \dots, X_n be a random sample of size n from the p.d.f.

$$f_{\theta}(x) = \theta(1 - \theta)^{x-1}, x = 1, 2, \dots,$$

$0 < \theta < 1$.

- Explain why $\sum_{i=1}^n X_i$ is a sufficient and complete statistic for $0 < \theta < 1$.

Henceforth let $n = 1$ and simply write $X_1 = X$. Define

$$T(X) = \begin{cases} 1, & \text{if } X = 1, \\ 0, & \text{if } X \geq 2. \end{cases}$$

- Show that T is the uniform minimum variance unbiased estimator of θ .
- Compute $E_{\theta}(X)$. (Hint: you may use the relation $\sum_{x=1}^{\infty} x(1 - \theta)^{x-1} = -\frac{d}{d\theta} \left\{ \sum_{x=1}^{\infty} (1 - \theta)^x \right\}, 0 < \theta < 1$.)
- Compute the Cramér-Rao lower bound for unbiased estimators of θ .
- Is T an efficient estimator of θ ? (Explain your answer.)

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7. Let X be a random sample of size 1 from the p.d.f.

$$f_{\theta}(x) = \left\{ 1 + \theta^2 \left(\frac{1}{2} - x \right) \right\} \mathbf{1}_{(0,1)}(x), x \in \mathbb{R},$$

$\theta \in [-1, 1]$. On the basis of this sample, consider the following.

- a. Derive the most powerful test for testing $H_0 : \theta = 0$ against $H_1 : \theta = \theta_1, \theta_1 \in [-1, 1], \theta_1 \neq 0$, at the level of significance $\alpha = 0.05$.
- b. Investigate whether or not the test derived in part a is uniformly most powerful at that level for testing $H_0 : \theta = 0$ against the alternative $\tilde{H}_1 : \theta \neq 0$.
- c. Determine the power function of the test in part a.

8. Let X_1, \dots, X_n be a random sample of size n from the p.d.f.

$$f_{\theta}(x) = \theta \cdot e^{-\theta x} \mathbf{1}_{(0,\infty)}(x), x \in \mathbb{R},$$

$\theta > 0$. Derive the family of likelihood ratio tests for testing $H_0 : \theta = \theta_0$, for fixed $\theta_0 > 0$, against the alternative $H_1 : \theta \neq \theta_0$, and prove that these tests are equivalent with rejecting H_0 when either $\bar{X} \leq c_1$ or $\bar{X} \geq c_2$ for suitable numbers $c_1 \leq c_2$.

9. Each person in a sample of size $n = 100$ from a population of males (M) and females (F) is classified according to whether he or she is a smoker (S) or not (NS). The data are summarized in the table below:

	S	NS	
M	20	35	55
F	15	30	45
	35	65	100

Test for independence of gender and the habit of smoking or not smoking at the approximate level of significance $\alpha = 0.05$.