

## Probability and Statistics Preliminary Examination: May 2024

### Instructions:

- Work all 5 Problems. Neither calculators nor electronic devices of any kind are allowed. Clearly state any theorem or fact that you use. Each of the 5 Problems is equally weighted (but the parts within a Problem may not be).
- Abbreviations/Acronyms.
  - pmf (probability mass function); pdf (probability density function); cdf (cumulative distribution function); mgf (moment generating function); iid (independent and identically distributed).
  - MSE (mean squared error), MOME (method of moments estimator); MLE (maximum likelihood estimator); PBE (posterior Bayes estimator); UMVUE (uniform minimum variance unbiased estimator); UMP (uniformly most powerful); LRT (likelihood ratio test).
- Notation.
  - $I(x \in A)$  or  $I_A(x)$ : indicator function for set  $A$ ; takes on the value 1 if  $x \in A$  and 0 otherwise.
  - $\mathbb{E}(X)$ : expectation of random variable  $X$ .
  - $\mathbb{V}(X)$ : variance of random variable  $X$ .
  - $X \sim N(a, b)$ :  $X$  has a normal distribution with mean  $a$  and variance  $b$ .
- Common distributions and other results.

**Bin**( $n, p$ ):  $\mathbb{E}(X) = np$ ,  $\mathbb{V}(X) = np(1 - p)$ , and pmf and mgf given respectively by

$$f(x) = \frac{n!}{x!(n-x)!} p^x (1-p)^{n-x} I(x \in \{0, 1, \dots, n\}), \quad M(t) = (1-p + pe^t)^n$$

**Poisson**( $\lambda$ ):  $\mathbb{E}(X) = \lambda$ ,  $\mathbb{V}(X) = \lambda$ , and pmf and mgf given respectively by

$$f(x) = \frac{e^{-\lambda} \lambda^x}{x!} I(x \in \{0, 1, \dots\}), \quad M(t) = \exp\{\lambda(e^t - 1)\}$$

**Geometric**( $p$ ):  $\mathbb{E}(X) = 1/p$ ,  $\mathbb{V}(X) = (1-p)/p^2$ , and pmf and mgf given respectively by

$$f(x) = p(1-p)^{x-1} I(x \in \{1, 2, \dots\}), \quad M(t) = \frac{pe^t}{1 - (1-p)e^t}, \quad t < -\log(1-p)$$

**Negative-Binomial**( $r, p$ ):  $\mathbb{E}(X) = r(1-p)/p$ ,  $\mathbb{V}(X) = r(1-p)/p^2$ , and pmf and mgf, respectively:

$$f(x) = \binom{r+x-1}{x} p^r (1-p)^x I(x \in \{0, 1, \dots\}), \quad M(t) = \left( \frac{p}{1 - (1-p)e^t} \right)^r, \quad t < -\log(1-p)$$

**Gamma**( $\alpha, \beta$ ):  $\mathbb{E}(X) = \alpha\beta$ ,  $\mathbb{V}(X) = \alpha\beta^2$ , and pdf and mgf given respectively by

$$f(x) = \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta} I(x > 0), \quad M(t) = \left( \frac{1}{1 - \beta t} \right)^\alpha, \quad t < 1/\beta$$

**Exp**( $\beta$ ): Gamma(1,  $\beta$ ).

**Beta**( $\alpha, \beta$ ):  $\mathbb{E}(X) = \frac{\alpha}{\alpha+\beta}$ ,  $\mathbb{V}(X) = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$ , and pdf given by

$$f(x) = \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1} I(0 < x < 1)$$

**Order Statistics:** Let  $X_{(1)} \leq \dots \leq X_{(n)}$  denote the order statistics from a random sample  $X_1, \dots, X_n$ . If  $X_1$  is continuous with pdf  $f(x)$  and cdf  $F(x)$ , the pdf of  $X_{(j)}$ ,  $(X_{(i)}, X_{(j)})$ , and  $(X_{(1)}, \dots, X_{(n)})$ , are given by:

$$f_{X_{(j)}}(x) = \frac{n!}{(j-1)!(n-j)!} [F(x)]^{j-1} [1-F(x)]^{n-j} f(x) I(-\infty < x < \infty)$$

$$f_{X_{(i)}, X_{(j)}}(x_i, x_j) = \frac{n!}{(i-1)!(j-1-i)!(n-j)!} f(x_i) f(x_j) [F(x_i)]^{i-1} [F(x_j) - F(x_i)]^{j-1-i} [1-F(x_j)]^{n-j} \\ \times I(-\infty < x_i \leq x_j < \infty)$$

$$f_{X_{(1)}, \dots, X_{(n)}}(x_1, \dots, x_n) = n! f(x_1) \cdots f(x_n) I(-\infty < x_1 \leq \dots \leq x_n < \infty)$$

1. Suppose  $X_1, \dots, X_n$  ( $n \geq 2$ ) are i.i.d from  $Unif(\theta, 2\theta)$  for some  $\theta > 0$ .
  - (a) Find the MLE of  $\theta$ .
  - (b) Is the MLE unbiased? Prove or disprove.

2. Let  $X, Y$  be two independent random variables following  $Unif(0, 1)$ . Find the density function of  $\frac{Y}{X+Y}$ .

3. Suppose  $X_1, \dots, X_n$  are i.i.d from  $Bin(1, \theta)$  for  $\theta \in (0, 1)$ . Let  $\bar{X}$  be the sample average.

(a) Find a function  $\tau$  such that  $\tau(\bar{X}) \xrightarrow{P} \theta(1 - \theta)$ .

(b) When  $\theta \neq 1/2$ , find the limiting distribution of  $\sqrt{n}(\tau(\bar{X}) - \theta(1 - \theta)) \xrightarrow{D}?$

(c) When  $\theta = 1/2$ , is the convergence in (b) still valid? Discuss the reason.

4. Suppose that  $X_1, \dots, X_n$  are i.i.d. observations from the cdf

$$F(x) = \begin{cases} 0 & \text{if } x \leq 0; \\ (x/\beta)^3 & \text{if } 0 < x \leq \beta; \\ 1 & \text{if } x > \beta. \end{cases}$$

- (a) Show that  $T = X_{(1)}/\beta$  is a pivotal quantity.
- (b) Based on  $T$ , find the value of  $u = u(X_1, \dots, X_n)$  such that  $[0, u]$  is a  $1 - \alpha$  confidence interval for  $\beta$ .

5. Suppose that  $X_1, \dots, X_n$  are i.i.d. observations from  $Beta(\theta, 1)$  with  $\theta > 0$ . Suppose that you only record the observations  $Y_i = I(X_i < 1/2)$  for  $i = 1, \dots, n$ .
- (a) Find the MLE of  $\theta$  based on the sample  $Y_1, \dots, Y_n$ .
  - (b) Let  $\hat{\theta}$  be the MLE in (a). Find the limiting distribution  $\sqrt{n}(\hat{\theta} - \theta) \xrightarrow{D} ?$
  - (c) Based on the result in (b), find an asymptotic  $1 - \alpha$  confidence interval for  $\theta$ .